

Lecture 5: The Goldreich-Levin Theorem

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In Lecture 3, we saw how a one-way permutation equipped with a hard-core predicate can be used to construct a pseudorandom generator. We now return to the question we left open there: do one-way functions have hard-core predicates?

Recap

In the last few lectures, we started to show how to construct a PRG from any OWP f .

1. We first showed that it suffices to construct a PRG with a single bit stretch, i.e. a PRG

$$G : \{0,1\}^\lambda \rightarrow \{0,1\}^{\lambda+1}.$$

This is the case since we showed that we can stretch any such PRG into one that has a polynomial stretch $k = \text{poly}(\lambda)$ by iterating G k times

$$G^k : \{0,1\}^\lambda \rightarrow \{0,1\}^{\lambda+k}.$$

2. We then defined the notion of a *hardcore predicate*

Definition 1. $P : \{0,1\}^\lambda \rightarrow \{0,1\}$ is a hardcore predicate of f if for every poly-size $\mathcal{A}$ there exists a negligible function μ such that for every $\lambda \in \mathbb{N}$,

$$\Pr_{x \leftarrow \{0,1\}^\lambda} [\mathcal{A}(f(x)) = P(x)] \leq \frac{1}{2} + \mu(\lambda).$$

3. We showed that if the OWP f had a hardcore predicate $P : \{0,1\}^\lambda \rightarrow \{0,1\}$ then the following function is a PRG:

$$G(x) = f(x) \circ P(x),$$

where $\circ$ denotes concatenation of strings. Thus, it remains to argue that every OWP has a hardcore predicate. Goldreich and Levin proved that every one-way function has a randomized hardcore predicate.

Goldreich-Levin Theorem

Theorem 2. *If f is a one-way function, then the following randomized predicate*

$$P(x, r) := x \cdot r \pmod{2} = \sum_{i \in \lambda} x_i r_i \pmod{2}$$

is a hardcore predicate for f . Namely, for every poly-size $\mathcal{A}$ there exists a negligible function $\mu : \mathbb{N} \rightarrow [0, 1]$ such that for every $\lambda \in \mathbb{N}$

$$\Pr_{x \leftarrow \{0,1\}^\lambda, r \leftarrow \{0,1\}^\lambda} [\mathcal{A}(f(x), r) = P(x, r)] \leq \frac{1}{2} + \mu(\lambda)$$

In particular, if f is a one-way permutation, then this theorem implies that

$$G(x, r) = f(x) \circ r \circ P(x, r)$$

is a PRG. In this lecture our focus is on proving this Theorem.

Proof. Suppose for contradiction that there exists a poly-size adversary $\mathcal{A}$ and a non-negligible function $\epsilon : \mathbb{N} \rightarrow [0, 1]$ such that for every $\lambda \in \mathbb{N}$,

$$\Pr_{x \leftarrow \{0,1\}^\lambda, r \leftarrow \{0,1\}^\lambda} [\mathcal{A}(f(x), r) = P(x, r)] \geq \frac{1}{2} + \epsilon(\lambda)$$

We will show how to use $\mathcal{A}$ to construct an adversary $\mathcal{B}$ that breaks the one-wayness of f .

Since $\mathcal{A}$ is only assumed to predict with small advantage ϵ , we cannot expect $\mathcal{B}$ to invert every given image $y = f(x)$. Indeed, it may be that $\mathcal{A}$ always fails when given an input y whose inverse starts with the bit 0. So which y 's can we hope to invert?

Define

$$\text{GOOD} = \{x \in \{0,1\}^\lambda : \Pr_r [\mathcal{A}(f(x), r) = x \cdot r] \geq \frac{1}{2} + \frac{\epsilon(\lambda)}{2}\}$$

By Markov's inequality

$$\Pr_{x \leftarrow \{0,1\}^\lambda} [x \in \text{GOOD}] \geq \frac{\epsilon(\lambda)}{2}$$

To expand on this, denoting by

$$p = \Pr_{x \leftarrow \{0,1\}^\lambda} [x \in \text{GOOD}],$$

it holds that

$$\begin{aligned}
\frac{1}{2} + \epsilon(\lambda) &\leq \\
\Pr_{x \leftarrow \{0,1\}^\lambda} [\mathcal{A}(f(x), r) = P(x, r)] &= \\
\Pr_{x \leftarrow \text{GOOD}} [\mathcal{A}(f(x), r) = P(x, r)] \cdot \Pr_{x \leftarrow \{0,1\}^\lambda} [x \in \text{GOOD}] &+ \Pr_{x \leftarrow \neg \text{GOOD}} [\mathcal{A}(f(x), r) = P(x, r)] \cdot \Pr_{x \leftarrow \{0,1\}^\lambda} [x \notin \text{GOOD}] = \\
\Pr_{x \leftarrow \text{GOOD}} [\mathcal{A}(f(x), r) = P(x, r)] \cdot p &+ \Pr_{x \leftarrow \neg \text{GOOD}} [\mathcal{A}(f(x), r) = P(x, r)] \cdot (1 - p) \leq \\
p + \frac{1}{2} + \frac{\epsilon(\lambda)}{2}, &
\end{aligned}$$

which implies that indeed $p \geq \frac{\epsilon(\lambda)}{2}$. This is usually called an averaging argument.

We will construct an algorithm $\mathcal{B}$ that succeeds in inverting every $f(x)$ for $x \in \text{GOOD}$ with non-negligible probability.

Special case 1: $\mathcal{A}$ predicts $P(x, r)$ perfectly for every $x \in \text{GOOD}$

As a warmup, suppose that $\mathcal{A}$ is a perfect predictor that for every $x \in \text{GOOD}$ and every $r \in \{0,1\}^\lambda$

$$\mathcal{A}(f(x), r) = x \cdot r$$

With such a perfect predictor it is easy to invert the OWF f . Specifically, the inverter $\mathcal{B}$, given $y = f(x)$ where $x \leftarrow \{0,1\}^\lambda$, does the following:

1. For every $i \in [\lambda]$ compute $x'_i = \mathcal{A}(y, e_i)$ where $e_i = (0, \dots, 0, 1, 0, \dots, 0)$ is the i -th standard basis vector with 1 in the i -th location and 0 everywhere else.
2. Output $x = (x'_1, \dots, x'_\lambda)$.

By our assumption that $\mathcal{A}$ is a perfect predictor, it follows that

$$x'_i = x \cdot e_i = x_i,$$

as desired.

Unfortunately, our predictor $\mathcal{A}$ may not be perfect. Next, we relax the perfect condition as follows.

Special case 2: There exists a non-negligible $\epsilon = \epsilon(\lambda)$ such that for every $x \in \text{GOOD}$, $\mathcal{A}$ predicts $P(x, r)$ with probability $3/4 + \epsilon$.

We can use this predictor to construct an algorithm $\mathcal{B}$ that for every $x \in \text{GOOD}$, given $f(x)$ finds an inverse with overwhelming probability (i.e., probability $1 - \mu$ for some negligible function μ). The inverter $\mathcal{B}$, given $y = f(x)$ where $x \leftarrow \{0,1\}^\lambda$, does the following:

1. Set $k = \Omega((\log \lambda)/\epsilon^2)$.
2. For every $i \in [\lambda]$, do the following:
 - (a) Choose at random $r_{i,1}, \dots, r_{i,k} \leftarrow \{0,1\}^\lambda$.
 - (b) For every $j \in [k]$ compute $b_{i,j} = \mathcal{A}(y, r_{i,j})$ and $b'_{i,j} = \mathcal{A}(y, r_{i,j} \oplus e_i)$, and let $x_{i,j} = b_{i,j} \oplus b'_{i,j}$.
 - (c) Let $x'_i = \text{majority}\{x_{i,1}, \dots, x_{i,k}\}$
3. Output $x' = (x'_1, \dots, x'_\lambda)$

Note that by our assumption for every $i \in [\lambda]$ and $j \in [k]$,

$$\begin{aligned}
 & \Pr_{r_{i,j} \leftarrow \{0,1\}^\lambda} [x_{i,j} = x_i] \geq \\
 & \Pr_{r_{i,j} \leftarrow \{0,1\}^\lambda} [\mathcal{A}(y, r_{i,j}) = x \cdot r_{i,j} \wedge \mathcal{A}(y, r_{i,j} \oplus e_i) = x \cdot (r_{i,j} \oplus e_i)] = \\
 & 1 - \Pr_{r_{i,j} \leftarrow \{0,1\}^\lambda} [\mathcal{A}(y, r_{i,j}) \neq x \cdot r_{i,j} \vee \mathcal{A}(y, r_{i,j} \oplus e_i) \neq x \cdot (r_{i,j} \oplus e_i)] \geq \\
 & 1 - \Pr_{r_{i,j} \leftarrow \{0,1\}^\lambda} [\mathcal{A}(y, r_{i,j}) \neq x \cdot r_{i,j}] - \Pr_{r_{i,j} \leftarrow \{0,1\}^\lambda} [\mathcal{A}(y, r_{i,j} \oplus e_i) \neq x \cdot (r_{i,j} \oplus e_i)] \geq \\
 & 1 - 1/2 + 2\epsilon = 1/2 + 2\epsilon
 \end{aligned}$$

Note that for every $i \in [\lambda]$ it holds that $x_{i,1}, \dots, x_{i,k}$ are independent Boolean random variables such that $\Pr[x_{i,j} = x_i] \geq 1/2 + 2\epsilon$. Therefore, by the Chernoff bound

$$\Pr[\text{majority}\{x'_{i,1}, \dots, x'_{i,k}\} \neq x_i] \leq 2^{-O(\epsilon^2 \cdot k)}$$

Theorem 3 (Chernoff bound). *Let $X_1, X_2, \dots, X_n$ be independent random variables taking values in $\{0,1\}$ (Bernoulli trials). Let*

$$X = \sum_{i=1}^n X_i, \quad \mu = \mathbb{E}[X].$$

For any $0 < \delta < 1$:

$$\Pr[X \geq (1 + \delta)\mu] \leq \exp\left(-\frac{\delta^2}{3}\mu\right),$$

and

$$\Pr[X \leq (1 - \delta)\mu] \leq \exp\left(-\frac{\delta^2}{2}\mu\right).$$

Alternatively, the additive version asserts that for every $t > 0$

$$\Pr[X > \mu + t] \leq \exp\left(-\frac{2t^2}{n}\right)$$

and

$$\Pr[X < \mu - t] \leq \exp\left(-\frac{2t^2}{n}\right)$$

Thus, for every $i \in [\lambda]$

$$\Pr[\mathcal{B}(f(x)) = x' : x'_i \neq x_i] \leq 1/\lambda^8,$$

which together with the union bound implies that

$$\Pr[\mathcal{B}(f(x)) \neq x] \leq \lambda \cdot 1/\lambda^8 = 1/\lambda^7.$$

Remark. Note that we could have chosen the same random $r_1, \dots, r_k \leftarrow \{0, 1\}^\lambda$ for all the coordinates $i \in [\lambda]$. The reason is that we do not need independence for the union bound.

General case: There exists a non-negligible $\epsilon = \epsilon(\lambda)$ such that for every $x \in \text{GOOD}$, $\mathcal{A}$ predicts $P(x, r)$ with probability $1/2 + \epsilon/2$.

Note that in the above analysis the reason that we needed a success probability of greater than $\frac{3}{4}$ is because, in order to get a non-negligible probability of guessing a single bit x_i we needed to combine *two* guesses for two bits $x \cdot r_{i,j}$ and $x \cdot (r_{i,j} \oplus e_i)$.

The Goldreich-Levin reduction starts with the following (ridiculous!) idea. Suppose that for each pair r and $r \oplus e_i$, we simply guess the bit $x \cdot r$ ourselves, and we only use the adversary $\mathcal{A}$ to guess the bit $x \cdot (r \oplus e_i)$. If we could somehow manage to always guess the bit $x \cdot r$ correctly then we could use the same argument as above since under this (ridiculous) assumption both bits $x \cdot r$ and $x \cdot (r \oplus e_i)$ are guessed correctly with probability at least $\frac{1}{2} + \epsilon/2$.

The problem is that we cannot guess $x \cdot r$ by ourselves with probability better than random. This is exactly the task that we needed $\mathcal{A}$ to do in the first place! But we can guess each of these bits at random and have a success probability $\frac{1}{2}$. This seems useless since in the above procedure, for each $i \in [\lambda]$ we needed to choose a total of k different values of $x \cdot r_j$, and if we guess each of them at random then we will guess all of them correctly only with probability 2^{-k} . But it turns out that this random guessing is actually really useful, due to the following clever idea: suppose that we happen to correctly guess $x \cdot s_1$ and $x \cdot s_2$. Then we can also guess correctly

$$x \cdot (s_1 \oplus s_2) = (x \cdot s_1) \oplus (x \cdot s_2)$$

More generally, suppose we guessed correctly

$$x \cdot r_1, x \cdot r_2, \dots, x \cdot r_\ell$$

then we can use these ℓ bits to guess correctly

$$x \cdot (\oplus_{i \in I} r_i) = \oplus_{i \in I} (x \cdot r_i)$$

for every $I \subseteq [\ell]$. Since there are 2^ℓ such subsets we learned 2^ℓ inner products $x \cdot r_I$, where $r_I = \oplus_{i \in I} r_i$. Therefore, in order to learn k

different inner products $x \cdot r_I$ we need to guess only $\ell = \log k$ many bits of the form $x \cdot r_j$! And we can guess all ℓ inner products correctly with probability $\frac{1}{k}$.

Still, all this may seem ridiculous since now these r_I 's are not independent! For example, $r_{\{1,2\}} = r_1 \oplus r_2$. As a result we cannot use the Chernoff bound in the analysis, so it seems that we fixed one problem but created another. However, the problem we created is not a real problem. While indeed the r_I 's are not all independent, they are *pairwise independent*, which turns out to be enough! Instead of using the Chernoff bound (which requires the random variables to be independent) we can rely on Chebyshev's inequality which is a weaker concentration bound but only requires the random variables to be pairwise independent.

Theorem 4 (Chebyshev's inequality). *Let $Z_1, \dots, Z_m$ be pairwise independent random variables obtaining values in $\{0, 1\}$, such that for every $j \in [m]$ $\Pr[Z_j = 1] = p$. Then, for every $\delta > 0$,*

$$\Pr \left[\left| \frac{\sum_{j=1}^m Z_j}{m} - p \right| \geq \delta \right] \leq \frac{1}{4\delta^2 m}$$

So, the inverter $\mathcal{B}$, given $y = f(x)$, does the following:

1. Set $k = \Omega(\lambda/\epsilon^2)$.
2. For every $i \in [\lambda]$, do the following:
 - (a) Set $\ell = \log(k+1)$
 - (b) choose at random $r_1, \dots, r_\ell \leftarrow \{0, 1\}^\lambda$.
 - (c) For every $j \in [\ell]$, choose at random $b_j \leftarrow \{0, 1\}$.¹
 - (d) For every $\emptyset \neq J \subseteq [\ell]$, let

¹ b_j is a guess for $x \cdot r_j$.

$$r_J = \oplus_{j \in J} r_j \quad \text{and} \quad b_J = \oplus_{j \in J} b_j.$$

- (e) Similarly, for every $\emptyset \neq J \subseteq [\ell]$, let

$$b'_{i,J} = \mathcal{A}(y, r_J \oplus e_i) \quad \text{and} \quad x_{i,J} = b_J \oplus b'_{i,J}.$$

- (f) Let $x'_i = \text{majority}\{x_{i,J}\}_{J \subseteq [\ell]}$

3. Output $x' = (x'_1, \dots, x'_\lambda)$.

Suppose that all our ℓ guesses were correct, which happens with probability $2^{-\ell} = \frac{1}{k}$. In this case, as above for every $i \in [\lambda]$ and

$J \subseteq [\ell]$,

$$\begin{aligned} & \Pr_{r_1, \dots, r_\ell \leftarrow \{0,1\}^\lambda} [x'_{i,J} = x_i] \geq \\ & \Pr_{r_1, \dots, r_\ell \leftarrow \{0,1\}^\lambda} [\mathcal{A}(y, (r_J + e_i)) = x \cdot (r_J \oplus e_i)] = \\ & 1 - \Pr_{r_1, \dots, r_\ell \leftarrow \{0,1\}^\lambda} [\mathcal{A}(y, r_J \oplus e_i) \neq x \cdot (r_J \oplus e_i)] \geq \\ & 1/2 + \epsilon/2. \end{aligned}$$

This seems great, except that now we cannot apply the Chernoff bound since the random variables $\{r_J\}_{J \subseteq [\ell]}$ are not independent! However, as mentioned above they are pairwise independent, and hence we can rely on Chebyshev's inequality to argue that for every $x \in \text{GOOD}$ and every $i \in [\lambda]$,

$$\Pr[x'_i \neq x_i] \leq \frac{1}{4\epsilon^2 k}$$

In particular, setting $k = \lambda \cdot \epsilon^{-2}$, we can take a union bound, and guarantee that for every $x \in \text{GOOD}$,

$$\Pr[\mathcal{B}(f(x)) = x] \geq \frac{1}{4},$$

contradicting the fact that f is a one-way function.

□

References